Graphical Models - Part I CMPT 419/726 Mo Chen

SFU Computing Science Feb. 10, 2020

Bishop PRML Ch. 8, some slides from Russell and Norvig AIMA2e

Outline

Probabilistic Models

Bayesian Networks

Markov Random Fields

Inference

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Probabilistic Models

- We now turn our focus to probabilistic models for pattern recognition
 - Probabilities express beliefs about uncertain events, useful for decision making, combining sources of information
- Key quantity in probabilistic reasoning is the joint distribution

$$p(x_1,x_2,\dots,x_K)$$

Where x_1 to x_K are all variables in model

- Address two problems
 - Inference: answering queries given the joint distribution
 - Learning: deciding what the joint distribution is (involves inference)
- All inference and learning problems involve manipulations of the joint distribution

Reminder - Three Tricks

· Bayes' rule:

$$p(Y|X) = \frac{p(X|Y)p(Y)}{p(X)} = \alpha p(X|Y)p(Y)$$

· Marginalization:

$$p(X) = \sum_{y} p(X, Y = y)$$
 or $p(X) = \int p(X, Y = y) dy$

· Product rule:

$$p(X,Y) = p(X)p(Y|X)$$

All 3 work with extra conditioning, e.g.:

$$p(X|Z) = \sum_{y} p(X, Y = y|Z)$$
$$p(Y|X, Z) = \alpha p(X|Y, Z)p(Y|Z)$$

Joint Distribution

	toothache		¬toothache	
	catch	$\neg catch$	catch	$\neg catch$
cavity	.108	.012	.072	.008
¬cavity	.016	.064	.144	.576

- Consider model with 3 boolean random variables: cavity, catch, toothache
- · Can answer query such as

 $p(\neg cavity | toothache)$

Joint Distribution

	toothache		¬toothache	
	catch	$\neg catch$	catch	$\neg catch$
cavity	.108	.012	.072	.008
¬cavity	.016	.064	.144	.576

- Consider model with 3 boolean random variables: cavity, catch, toothache
- Can answer query such as

$$p(\neg cavity | toothache) = \frac{p(\neg cavity, toothache)}{p(toothache)}$$
$$p(\neg cavity | toothache) = \frac{0.016 + 0.064}{0.108 + 0.012 + 0.016 + 0.064} = 0.4$$

Joint Distribution

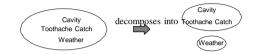
• In general, to answer a query on random variables $\mathbf{Q} = Q_1, \dots, Q_N$ given evidence $\mathbf{E} = \mathbf{e}, \mathbf{E} = E_1, \dots, E_M, \mathbf{e} = e_1, \dots, e_M$:

$$p(\mathbf{Q}|\mathbf{E}=e) = \frac{p(\mathbf{Q}, \mathbf{E}=e)}{P(\mathbf{E}=e)}$$
$$= \frac{\sum_{h} p(\mathbf{Q}, \mathbf{E}=e, \mathbf{H}=h)}{\sum_{q,h} p(\mathbf{Q}=q, \mathbf{E}=e, \mathbf{H}=h)}$$

Problems

- · The joint distribution is large
 - e. g. with K boolean random variables, 2^K entries
- Inference is slow, previous summations take $O(2^K)$ time
- Learning is difficult, data for 2^K parameters
- · Analogous problems for continuous random variables

Reminder - Independence



- A and B are independent iff p(A|B) = p(A) or p(B|A) = p(B) or p(A,B) = p(A)p(B)
- p(Toothache, Catch, Cavity, Weather) =
 p(Toothache, Catch, Cavity)p(Weather)
 - 32 entries reduced to 12 (Weather takes one of 4 values)
- Absolute independence powerful but rare
- Dentistry is a large field with hundreds of variables, none of which are independent. What to do?

Reminder - Conditional Independence

- p(Toothache, Cavity, Catch) has $2^3 1 = 7$ independent entries
- If I have a cavity, the probability that the probe catches in it doesn't depend on whether I have a toothache:

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P(catch|toothache, cavity) = P(catch|cavity)
```

- The same independence holds if I haven't got a cavity: $P(catch|toothache, \neg cavity) = P(catch|\neg cavity)$
- Catch is conditionally independent of Toothache given Cavity: p(Catch|Toothache, Cavity) = p(Catch|Cavity)
- Equivalent statements:
 - p(Toothache|Catch, Cavity) = p(Toothache|Cavity)
 - p(Toothache, Catch|Cavity) = p(Toothache|Cavity)p(Catch|Cavity)

Conditional Independence contd.

- Write out full joint distribution using chain rule:
 p(Toothache, Catch, Cavity)
 = p(Toothache|Catch, Cavity)p(Catch, Cavity)
 = p(Toothache|Catch, Cavity)p(Catch|Cavity)p(Cavity)
 = p(Toothache|Cavity)p(Catch|Cavity)p(Cavity)
 2 + 2 + 1 = 5 independent numbers
- In many cases, the use of conditional independence greatly reduces the size of the representation of the joint distribution

Graphical Models

- Graphical Models provide a visual depiction of probabilistic model
- · Conditional indepence assumptions can be seen in graph
- Inference and learning algorithms can be expressed in terms of graph operations
- We will look at 2 types of graph (can be combined)
 - Directed graphs: Bayesian networks
 - Undirected graphs: Markov Random Fields
 - Factor graphs (won't cover)

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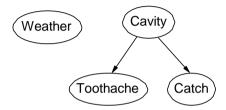
Bayesian Networks

- A simple, graphical notation for conditional independence assertions and hence for compact specification of full joint distributions
- Syntax:
 - a set of nodes, one per variable
 - a directed, acyclic graph (link ≈ "directly influences")
 - a conditional distribution for each node given its parents:

$$p(X_i|pa(X_i))$$

 In the simplest case, conditional distribution represented as a conditional probability table (CPT) giving the distribution over X_i for each combination of parent values

Example

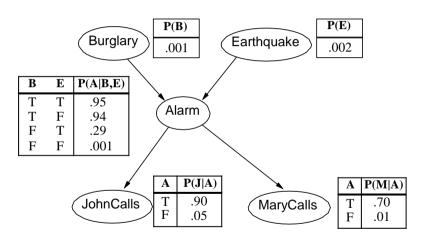


- Topology of network encodes conditional independence assertions:
 - Weather is independent of the other variables
 - Toothache and Catch are conditionally independent given Cavity

Example

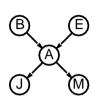
- I'm at work, neighbor John calls to say my alarm is ringing, but neighbor Mary doesn't call. Sometimes it's set off by minor earthquakes. Is there a burglar?
- Variables: Burglar, Earthquake, Alarm, JohnCalls, MaryCalls
- Network topology reflects "causal" knowledge:
 - A burglar can set the alarm off
 - An earthquake can set the alarm off
 - · The alarm can cause Mary to call
 - · The alarm can cause John to call

Example contd.



Compactness

- A CPT for Boolean X_i with k Boolean parents Has 2^k rows for the combinations of parent values
- Each row requires one number p for $X_i = true$ (the number for $X_i = false$ is just 1 p)
- If each variable has no more than k parents, the complete network requires $O(n \cdot 2^k)$ numbers
- i.e., grows linearly with n, vs. $O(2^n)$ for the full joint distribution
- For burglary net, ?? numbers
 - 1 + 1 + 4 + 2 + 2 = 10 numbers (vs. $2^5 - 1 = 31$)



Global Semantics

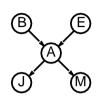
 Global semantics defines the full joint distribution as the product of the local conditional distributions:

$$P(x_1, \dots, x_n) = \prod_{i=1}^n p(x_i|pa(X_i))$$

e.g.
$$P(j \land m \land a \land \neg b \land \neg e) =$$

$$P(j|a)P(m|a)P(a|\neg b, \neg e)P(\neg b)P(\neg e)$$

= 0.9 × 0.7 × 0.001 × 0.999 × 0.998
≈ 0.00063



Constructing Bayesian Networks

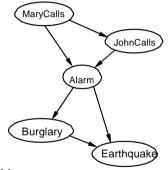
- Need a method such that a series of locally testable assertions of conditional independence guarantees the required global semantics
 - 1. Choose an ordering of variables $X_1, ..., X_n$
 - 2. For i= 1 to n add X_i to the network select parents from X_1, \ldots, X_{i-1} such that $p(X_i|pa(X_i)) = p(X_i|X_1, \ldots, X_{i-1})$

This choice of parents guarantees the global semantics:

$$\begin{split} p(X_1,\ldots,X_n) &= \prod_{i=1}^n p(X_i|X_1,\ldots,X_{i-1}) & \text{(chain rule)} \\ &= \prod_i p\big(X_i|pa(X_i)\big) & \text{(by construction)} \end{split}$$

Example

Suppose we choose the ordering M, J, A, B, E



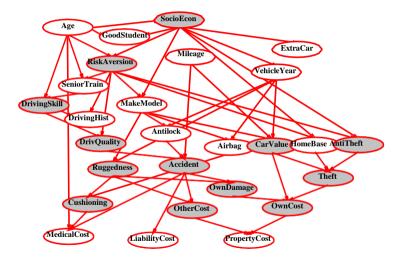
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P(J|M) = P(J)? No P(A|J,M) = P(A|J)? P(A|J,M) = P(A)? No P(B|A,J,M) = P(B|A)? Yes P(B|A,J,M) = P(B)? No P(E|B,A,J,M) = P(E|A)? No P(E|B,A,J,M) = P(E|A,B)? Yes
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Example contd.

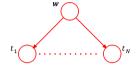


- Deciding conditional independence is hard in noncausal directions
 - (Causal models and conditional independence seem hardwired for humans!)
- Assessing conditional probabilities is hard in noncausal directions
- Network is less compact: 1 + 2 + 4 + 2 + 4 = 13 numbers needed

Example - Car Insurance



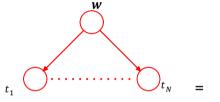
Example - Polynomial Regression

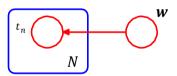


- · Bayesian polynomial regression model
- Observations $\mathbf{t} = (t_1, ..., t_N)$
- Vector of coefficients w
- Inputs x and noise variance σ^2 were assumed fixed, not stochastic and hence not in model
- Joint distribution:

$$p(t, w) = p(w) \prod_{n=1}^{N} p(t_n | w)$$

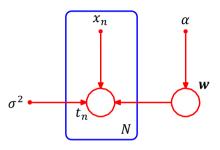
Plates





- A shorthand for writing repeated nodes such as the t_n uses plates

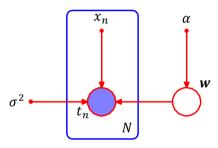
Deterministic Model Parameters



- Can also include deterministic parameters (not stochastic) as small nodes
- Bayesian polynomial regression model:

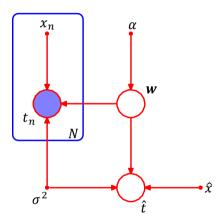
$$p(\boldsymbol{t}, \boldsymbol{w} | \boldsymbol{x}, \alpha, \sigma^2) = p(\boldsymbol{w} | \alpha) \prod_{n=1}^{N} p(t_n | \boldsymbol{w}, x_n, \sigma^2)$$

Observations



- In polynomial regression, we assumed we had a training set of N pairs (x_n, t_n)
- Convention is to use shaded nodes for observed random variables

Predictions



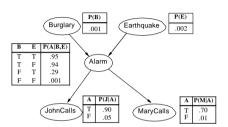
- Suppose we wished to predict the value \hat{t} for a new input \hat{x}
- The Bayesian network used for this inference task would be this one

Specifying Distributions - Discrete Variables

- Earlier we saw the use of conditional probability tables (CPT) for specifying a distribution over discrete random variables with discrete-valued parents
- For a variable with no parents, with K possible states:

$$p(\boldsymbol{x}|\boldsymbol{\mu}) = \prod_{k=1}^K \mu_k^{x_k}$$

• e.g. $p(B) = 0.001^{B_1}0.999^{B_2}$, 1-of-K representation



Specifying Distributions - Discrete Variables cont.

• With two variables x_1, x_2 can have two cases



Dependent

$$p(\mathbf{x}_{1}, \mathbf{x}_{2} | \boldsymbol{\mu}) = p(\mathbf{x}_{1} | \boldsymbol{\mu}) p(\mathbf{x}_{2} | \mathbf{x}_{1}, \boldsymbol{\mu}) \quad p(\mathbf{x}_{1}, \mathbf{x}_{2} | \boldsymbol{\mu}) = p(\mathbf{x}_{1} | \boldsymbol{\mu}) p(\mathbf{x}_{2} | \boldsymbol{\mu})$$

$$= \left(\prod_{k=1}^{K} \mu_{k1}^{x_{1k}} \right) \left(\prod_{k=1}^{K} \prod_{j=1}^{K} \mu_{kj2}^{x_{1k}x_{2j}} \right) \quad = \left(\prod_{k=1}^{K} \mu_{k1}^{x_{1k}} \right) \left(\prod_{k=1}^{K} \mu_{k2}^{x_{2k}} \right)$$

• $K^2 - 1$ free parameters in μ



Independent

$$p(\mathbf{x}_{1}, \mathbf{x}_{2} | \boldsymbol{\mu}) = p(\mathbf{x}_{1} | \boldsymbol{\mu}) p(\mathbf{x}_{2} | \boldsymbol{\mu})$$
$$= \left(\prod_{k=1}^{K} \mu_{k1}^{x_{1k}} \right) \left(\prod_{k=1}^{K} \mu_{k2}^{x_{2k}} \right)$$

• 2(K-1) free parameters in μ

Chains of Nodes



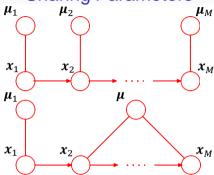
- With M nodes, could form a chain as shown above
- Number of parameters is:

$$(K-1) + (M-1)K(K-1)$$

others

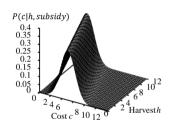
- Compare to:
 - K^M − 1 for fully connected graph
 - M(K-1) for graph with no edges (all independent)

Sharing Parameters



- Another way to reduce number of parameters is sharing parameters (a. k. a. tying of parameters)
- Lower graph reuses same μ for nodes 2 M
 - μ is a random variable in this network, could also be deterministic
- (K-1) + K(K-1) parameters

Specifying Distributions - Continuous Variables



 One common type of conditional distribution for continuous variables is the linear-Gaussian

$$p(x_i|pa_i) = \mathcal{N}\left(x_i; \sum_{j \in pa_i} w_{ij}x_j + b_i, v_i\right)$$

e.g. With one parent Harvest:

$$p(c|h) = \mathcal{N}(c; -0.5h + 5, 1)$$

For harvest h, mean cost is
 -0.5h + 5, variance is 1

Linear Gaussian

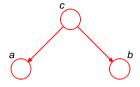
 Interesting fact: if all nodes in a Bayesian Network are linear Gaussian, joint distribution is a multivariate Gaussian

$$p(x_i|pa_i) = \mathcal{N}\left(x_i; \sum_{j \in pa_i} x_{ij}x_j + b_i, v_i\right)$$
$$p(x_1, \dots, x_N) = \prod_{i=1}^N \mathcal{N}\left(x_i; \sum_{j \in pa_i} x_{ij}x_j + b_i, v_i\right)$$

- Each factor looks like $\exp\Big(\big(x_i \pmb{w}_i^{\mathsf{T}} \pmb{x}_{pa_i}\big)^2\Big)$, this product will be another quadratic form
- With no links in graph, end up with diagonal covariance matrix
- With fully connected graph, end up with full covariance matrix

Conditional Independence in Bayesian Networks

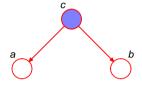
- Recall again that a and b are conditionally independent given c ($a \perp b \mid c$) if
 - p(a|b,c) = p(a|c) or equivalently
 - p(a,b|c) = p(a|c)p(b|c)
- Before we stated that links in a graph are ≈ "directly influences"
- We now develop a correct notion of links, in terms of the conditional independences they represent
 - This will be useful for general-purpose inference methods



The graph above means

$$p(a,b,c) = p(a|c)p(b|c)p(c)$$
$$p(a,b) = \sum_{c} p(a|c)p(b|c)p(c)$$
$$\neq p(a)p(b) \text{ in general}$$

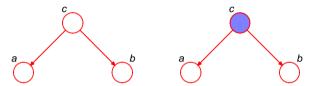
So a and b not independent



However, conditioned on c,

$$p(a,b|c) = \frac{p(a,b,c)}{p(c)} = \frac{p(a|c)p(b|c)p(c)}{p(c)} = p(a|c)p(b|c)$$

So a ⊥ b|c



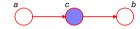
- Note the path from a to b in the graph
 - When c is not observed, path is open, a and b not independent
 - When c is observed, path is blocked, a and b independent
- In this case c is tail-to-tail with respect to this path



The graph above means

$$p(a,b,c) = p(a)p(b|c)p(c|a)$$

Again a and b not independent



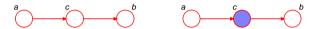
However, conditioned on c

$$p(a,b|c) = \frac{p(a,b,c)}{p(c)} = \frac{p(a)p(b|c)}{p(c)}p(c|a)$$

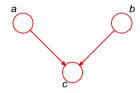
$$= \frac{p(a)p(b|c)}{p(c)} \underbrace{\frac{p(a|c)p(c)}{p(a)}}_{\text{Bayes' rule}}$$

$$= p(a|c)p(b|c)$$

So a ⊥ b|c



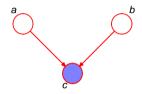
- As before, the path from a to b in the graph
 - When c is not observed, path is open, a and b not independent
 - When c is observed, path is blocked, a and b independent
- In this case c is head-to-tail with respect to this path



The graph above means

$$p(a,b,c) = p(a)p(b)p(c|a,b)$$
$$p(a,b) = \sum_{c} p(a)p(b)p(c|a,b)$$
$$= p(a)p(b)$$

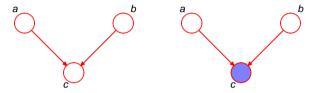
This time a and b are independent



However, conditioned on c

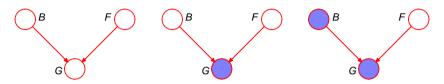
$$p(a,b|c) = \frac{p(a,b,c)}{p(c)} = \frac{p(a)p(b)p(c|a,b)}{p(c)}$$

$$\neq p(a|c)p(b|c) \text{ in general}$$



- · Frustratingly, the behaviour here is different
 - When c is not observed, path is blocked, a and b independent
 - When c is observed, path is unblocked, a and b not independent
- In this case c is head-to-head with respect to this path
- Situation is in fact more complex, path is unblocked if any descendent of c is observed

Part 3 - Intuition

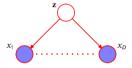


- Binary random variables B (battery charged), F (fuel tank full), G (fuel gauge reads full)
- B and F independent
- But if we observe G = 0 (false) things change
 - e.g. p(F=0|G=0,B=0) could be less than p(F=0|G=0), as B=0 explains away the fact that the gauge reads empty
 - Recall that p(F|G,B) = p(F|G) is another $F \perp \!\!\! \perp B|G$

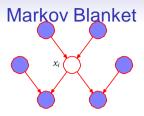
D-separation

- · A general statement of conditional independence
- For sets of nodes A, B, C, check all paths from A to B in graph
- If all paths are blocked, then $A \perp \!\!\!\perp B \mid C$
- Path is blocked if:
 - Arrows meet head-to-tail or tail-to-tail at a node in C
 - Arrows meet head-to-head at a node, and neither node nor any descendent is in C

Naive Bayes



- · Commonly used naive Bayes classification model
- Class label z, features x₁, ..., x_D
- Model assumes features independent given class label
 - Tail-to-tail at z, blocks path between features



- What is the minimal set of nodes which makes a node x_i conditionally independent from the rest of the graph?
 - x_i 's parents, children, and children's parents (co-parents)
- Define this set MB, and consider:

$$p(x_i|x_{\{j\neq i\}}) = \frac{p(x_1, \dots, x_D)}{\int p(x_1, \dots, x_D) dx_i}$$
$$= \frac{\prod_k p(x_k|pa_k)}{\int \prod_k p(x_k|pa_k) dx_i}$$

• All factors other than those for which x_i is x_k or in pa_k cancel



Learning Parameters

- When all random variables are observed in training data, relatively straight-forward
 - Distribution factors, all factors observed
 - e.g. Maximum likelihood used to set parameters of each Distribution $p(x_i|pa_i)$ separately
- When some random variables not observed, it's tricky
 - This is a common case
 - Expectation-maximization (later) is a method for this

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